### Introduction to stochastic processes

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## A quick introduction to probability theory

## **Probability space** $(\Omega, \mathcal{A}, \mathbb{P})$

• Goal: Model a random experiment.

Example 1: toss a coin.

Example 2: pick a number between 0 and 1 (matlab function rand).

•  $\Omega$ : fundamental set (set of all possible realisations).

 $\hookrightarrow$  A realization  $\omega$  is an element of  $\Omega$ .

Example 1:  $\Omega = \{H, T\}$ . Example 2:  $\Omega = [0, 1]$ .

•  $\mathcal{A}$ :  $\sigma$ -algebra on  $\Omega$  (set of events).

Example 1:  $\mathcal{A} = \mathcal{P}(\Omega);$ 

an event: the result is head,  $A = \{H\}$ .

Example 2:  $\mathcal{A} = \mathcal{B}([0,1]);$ 

an event: the number is larger than 1/2, A = [1/2, 1].

P: probability (gives the probability of events): function from A to [0,1] such that:
P(Ω) = 1,

- if  $(A_j)_{j\geq 1}$  is a numerable family of disjoint sets of  $\mathcal{A}$  then  $\mathbb{P}(\cup_j A_j) = \sum_j \mathbb{P}(A_j)$ . Example 1:  $\mathbb{P}(\{H\}) = \mathbb{P}(\{T\}) = \frac{1}{2}$ . Example 2:  $\mathbb{P}([a,b]) = b - a$  for all  $0 \leq a \leq b \leq 1$ .

ESNT

## **Random variable**

• Random variable X = random number.

Application  $X : \Omega \to \mathbb{R}$ .

A realization  $X(\omega)$  of a random variable is a real number.

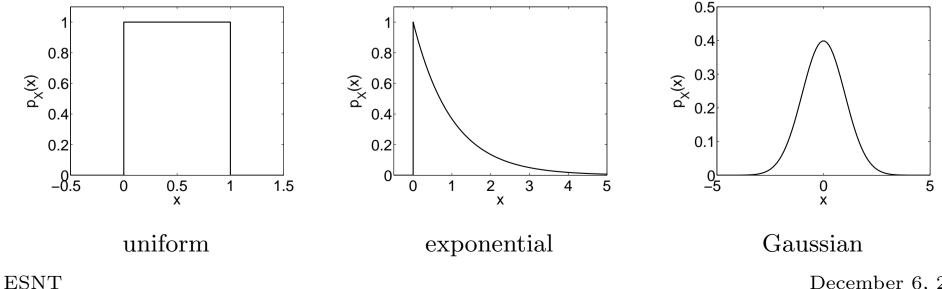
• The distribution of a random variable is characterized by moments of the form  $\mathbb{E}[\phi(X)]$  for  $\phi \in \mathcal{C}_b(\mathbb{R}, \mathbb{R})$ : 0

$$\mathbb{E}[\phi(X)] = \int_{\Omega} \phi(X(\omega)) \mathbb{P}(d\omega)$$

• The distribution of a (continuous) random variable is characterized by the probability density function (pdf)  $p_X$ :

$$\mathbb{E}[\phi(X)] = \int_{-\infty}^{\infty} \phi(x) p_X(x) dx$$

• Usual pdfs:



• The mean (expectation) of a random variable X with pdf  $p_X(x)$  is

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} x p_X(x) dx$$

The variance of a random variable X with pdf p(x) is

$$\operatorname{Var}(X) = \mathbb{E}\left[(X - \mathbb{E}[X])^2\right] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \int_{-\infty}^{\infty} x^2 p_X(x) dx - \left(\int_{-\infty}^{\infty} x p_X(x) dx\right)^2$$

The variance measures the dispersion of the random variable (around its mean).

• A standard Gaussian random variable X has the pdf

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

Its mean is  $\mathbb{E}[X] = 0$  and its variance is  $\operatorname{Var}(X) = 1$ . We write  $X \sim \mathcal{N}(0, 1)$ .

• A Gaussian random variable X with mean  $\mu$  and variance  $\sigma^2$  has the pdf

$$p_X(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

We write  $X \sim \mathcal{N}(\mu, \sigma^2)$ .

#### ESNT

### **Random vector**

• *n*-dimensional random vector  $\mathbf{X}$  = collection of *n* random variables  $(X_1, \ldots, X_n)$ . Application  $\mathbf{X} : \Omega \to \mathbb{R}^n$ .

A realization  $X(\omega)$  of a random vector is a vector in  $\mathbb{R}^n$ .

The distribution of a (continuous) random vector is characterized by the pdf  $p_{\mathbf{X}}$ :

$$\mathbb{E}[\phi(\boldsymbol{X})] = \int_{\mathbb{R}^n} \phi(\boldsymbol{x}) p_{\boldsymbol{X}}(\boldsymbol{x}) d\boldsymbol{x}, \qquad \forall \phi \in \mathcal{C}_b(\mathbb{R}^n, \mathbb{R})$$

The vector  $\boldsymbol{X} = (X_1, \ldots, X_n)$  is independent if

$$p_{\boldsymbol{X}}(\boldsymbol{x}) = \prod_{j=1}^{n} p_{X_j}(x_j)$$

or equivalently

$$\mathbb{E}[\phi_1(X_1)\cdots\phi_n(X_n)] = \mathbb{E}[\phi_1(X_1)]\cdots\mathbb{E}[\phi_n(X_n)], \qquad \forall \phi_1,\ldots,\phi_n \in \mathcal{C}_b(\mathbb{R},\mathbb{R})$$

Example: a normalized Gaussian random vector  $\boldsymbol{X}$  has the Gaussian pdf

$$p_{\boldsymbol{X}}(\boldsymbol{x}) = \frac{1}{\sqrt{(2\pi)^n}} \exp\left(-\frac{|\boldsymbol{x}|^2}{2}\right)$$

It is a vector of independent random normalized Gaussian variables.

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#### Limit theorems

#### • Law of Large Numbers.

Let  $(X_n)_{n\geq 0}$  be independent and identically distributed (i.i.d.) random variables. If  $\mathbb{E}[|X_1|] < \infty$ , then

$$\bar{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n) \xrightarrow{n \to \infty} m$$
 almost surely, with  $m = \mathbb{E}[X_1]$ 

"The empirical mean converges to the statistical mean".

## • Central Limit Theorem. Fluctuations theory. Let $(X_n)_{n\geq 0}$ be i.i.d. random variables. If $\mathbb{E}[X_1^2] < \infty$ , then

$$\sqrt{n} \left( \bar{X}_n - m \right) = \sqrt{n} \left( \frac{1}{n} (X_1 + X_2 + \dots + X_n) - m \right) \xrightarrow{n \to \infty} \mathcal{N}(0, \sigma^2) \text{ in law}$$
where
$$\begin{cases}
m = \mathbb{E}[X_1] \\
\sigma^2 = \mathbb{E}[X_1^2] - \mathbb{E}[X_1]^2 = \mathbb{E}[(X_1 - \mathbb{E}[X_1])^2]
\end{cases}$$

"For large n, the error  $\bar{X}_n - m$  obeys the Gaussian distribution  $\mathcal{N}(0, \sigma^2/n)$ ."

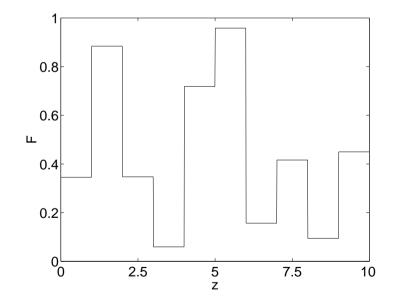
**Stochastic** processes

## Toy model

Let  $F(z) \in \mathbb{R}$  be the stepwise constant random process

$$F(z) = \sum_{i=1}^{\infty} F_i \mathbf{1}_{[i-1,i)}(z)$$

where  $F_i$  independent random variables  $\mathbb{E}[F_i] = \overline{F}$  and  $\mathbb{E}[(F_i - \overline{F})^2] = \sigma^2$ .



Here  $F_i$  are independent with uniform distribution over (0, 1).

### **Random process**

• Random variable X = random number.

A realization of the random variable = a real number.

Distribution of X characterized by moments of the form  $\mathbb{E}[\phi(X)]$  where  $\phi \in \mathcal{C}_b(\mathbb{R}, \mathbb{R})$ .

Stochastic process (μ(x))<sub>x∈R<sup>d</sup></sub> = random function.
A realization of the process = a function from R<sup>d</sup> to R.
Distribution of (μ(x))<sub>x∈R<sup>d</sup></sub> characterized by moments of the form
E[φ(μ(x<sub>1</sub>),...,μ(x<sub>n</sub>))], for any n ∈ N\*, x<sub>1</sub>,..., x<sub>n</sub> ∈ R<sup>d</sup>, φ ∈ C<sub>b</sub>(R<sup>n</sup>, R).

Example: Gaussian process.

#### Gaussian process

• Gaussian process  $(\mu(\boldsymbol{x}))_{\boldsymbol{x}\in\mathbb{R}^d}$  characterized by its first two moments  $m(\boldsymbol{x}_1) = \mathbb{E}[\mu(\boldsymbol{x}_1)]$  and  $R(\boldsymbol{x}_1, \boldsymbol{x}_2) = \mathbb{E}[\mu(\boldsymbol{x}_1)\mu(\boldsymbol{x}_2)].$ 

Any linear combination  $\mu_{\lambda} = \sum_{i=1}^{n} \lambda_i \mu(\boldsymbol{x}_i)$  has Gaussian distribution  $\mathcal{N}(m_{\lambda}, \sigma_{\lambda}^2)$  with

$$m_{\lambda} = \sum_{i=1}^{n} \lambda_i \mathbb{E}[\mu(\boldsymbol{x}_i)] \text{ and } \sigma_{\lambda}^2 = \sum_{i,j=1}^{n} \lambda_i \lambda_j \mathbb{E}[\mu(\boldsymbol{x}_i)\mu(\boldsymbol{x}_j)] - m_{\lambda}^2$$

• Simulation: in order to simulate  $(\mu(\boldsymbol{x}_1), \ldots, \mu(\boldsymbol{x}_n))$ :

- evaluate the mean vector  $M_i = \mathbb{E}[\mu(\boldsymbol{x}_i)]$  and the covariance matrix  $C_{ij} = \mathbb{E}[\mu(\boldsymbol{x}_i)\mu(\boldsymbol{x}_j)] - \mathbb{E}[\mu(\boldsymbol{x}_i)]\mathbb{E}[\mu(\boldsymbol{x}_j)].$ 

- generate a random vector  $\mathbf{X} = (X_1, \ldots, X_n)$  of *n* independent Gaussian random variables with mean 0 and variance 1.

- compute  $\boldsymbol{Y} = \boldsymbol{M} + \mathbf{C}^{1/2} \boldsymbol{X}$ . The vector  $\boldsymbol{Y}$  has the distribution of  $(\mu(\boldsymbol{x}_1), \dots, \mu(\boldsymbol{x}_n))$ . Note: the computation of the square root is expensive (use Cholesky method).

### **Brownian motion**

• Brownian motion  $(W_z)_{z\geq 0}$  (starting from 0)= real Gaussian process with mean 0 and covariance function

$$\mathbb{E}[W_z W_{z'}] = z \wedge z'$$

The realizations of the Brownian motion are continuous but not differentiable.

The increments of the Brownian motion are independent:

if  $z_n \ge z_{n-1} \ge \cdots \ge z_1 \ge z_0 = 0$ , then  $(W_{z_n} - W_{z_{n-1}}, \dots, W_{z_2} - W_{z_1}, W_{z_1})$  are independent Gaussian random variables with mean 0 and variance

$$\mathbb{E}[(W_{z_j} - W_{z_{j-1}})^2] = z_j - z_{j-1}$$

• Simulation: in order to simulate  $(W_h, W_{2h}, ..., W_{nh})$ :

- evaluate the covariance matrix  $\mathbf{C} = (C_{jl})_{j=,l=1,\ldots,n}$  with  $C_{jl} = (j \wedge l)h$ .

- generate a random vector  $\mathbf{X} = (X_1, ..., X_n)$  of *n* independent Gaussian random variables with mean 0 and variance 1.

- compute  $\boldsymbol{Y} = \mathbf{C}^{1/2} \boldsymbol{X}$ .

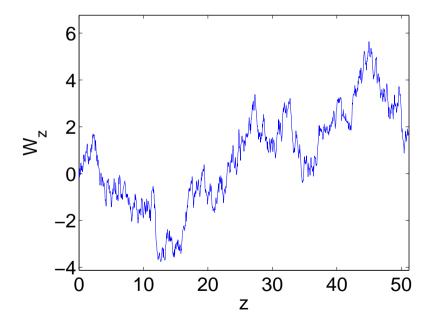
The vector  $\boldsymbol{Y}$  has the distribution of  $(W_h, W_{2h}, ..., W_{nh})$ .

• Simulation: in order to simulate  $(W_h, W_{2h}, \ldots, W_{nh})$ :

- generate a random vector  $\mathbf{X} = (X_1, \dots, X_n)^T$  of *n* independent Gaussian random variables with mean 0 and variance 1.

- compute  $Y_j = \sqrt{h} \sum_{i=1}^j X_i$ .

The vector  $\boldsymbol{Y}$  has the distribution of  $(W_h, W_{2h}, \ldots, W_{nh})^T$ .



### **Stationary random process**

•  $(\mu(\boldsymbol{x}))_{\boldsymbol{x}\in\mathbb{R}^d}$  is stationary if  $(\mu(\boldsymbol{x}+\boldsymbol{x}_0))_{\boldsymbol{x}\in\mathbb{R}^d}$  has the same distribution as  $(\mu(\boldsymbol{x}))_{\boldsymbol{x}\in\mathbb{R}^d}$  for any  $\boldsymbol{x}_0\in\mathbb{R}^d$ . Sufficient and necessary condition:

$$\mathbb{E}\big[\phi(\mu(\boldsymbol{x}_1),\ldots,\mu(\boldsymbol{x}_n))\big] = \mathbb{E}\big[\phi(\mu(\boldsymbol{x}_0+\boldsymbol{x}_1),\ldots,\mu(\boldsymbol{x}_0+\boldsymbol{x}_n))\big]$$

for any  $n, \boldsymbol{x}_0, \ldots, \boldsymbol{x}_n \in \mathbb{R}^d, \phi \in \mathcal{C}_b(\mathbb{R}^n, \mathbb{R}).$ 

• Example: Gaussian process  $\mu(\boldsymbol{x})$  with mean zero  $\mathbb{E}[\mu(\boldsymbol{x})] = 0 \ \forall \boldsymbol{x}$  and covariance function  $\mathbb{E}[\mu(\boldsymbol{x}')\mu(\boldsymbol{x}'+\boldsymbol{x})] = c(\boldsymbol{x}).$ 

• Bochner's theorem: a function  $c(\mathbf{x})$  is a covariance function of a stationary process if and only if its Fourier transform  $\hat{c}(\mathbf{k})$  is nonnegative.

$$\hat{c}(oldsymbol{k}) = \int_{\mathbb{R}^d} e^{ioldsymbol{k}\cdotoldsymbol{x}} c(oldsymbol{x}) doldsymbol{x}$$

• Spectral representation (of real-valued stationary Gaussian process):

$$\mu(\boldsymbol{x}) = \frac{1}{(2\pi)^d} \int_{\mathbb{R}^d} e^{-i\boldsymbol{k}\cdot\boldsymbol{x}} \sqrt{\hat{c}(\boldsymbol{k})} \hat{n}_{\boldsymbol{k}} d\boldsymbol{k}$$

with  $\hat{n}_{k}$  complex white noise, i.e.:

 $\hat{n}_{k}$  complex-valued, Gaussian,  $\hat{n}_{-k} = \overline{\hat{n}_{k}}$ ,  $\mathbb{E}[\hat{n}_{k}] = 0$ , and  $\mathbb{E}[\hat{n}_{k}\overline{\hat{n}_{k'}}] = (2\pi)^{d}\delta(k - k')$ . (the representation is formal, one should use stochastic integrals  $d\hat{W}_{k} = \hat{n}_{k}dk$ ).

We have  $\hat{n}_{\mathbf{k}} = \int e^{i\mathbf{k}\cdot\mathbf{x}}n(\mathbf{x})d\mathbf{x}$  where  $n(\mathbf{x})$  is a real white noise, i.e.:  $n(\mathbf{x})$  real-valued, Gaussian,  $\mathbb{E}[n(\mathbf{x})] = 0$ , and  $\mathbb{E}[n(\mathbf{x})n(\mathbf{x}')] = \delta(\mathbf{x} - \mathbf{x}')$ . (in 1D, formally,  $n(x) = dW_x/dx$ ). • Spectral representation (of real-valued stationary Gaussian process):

$$\mu(\boldsymbol{x}) = \frac{1}{(2\pi)^d} \int_{\mathbb{R}^d} e^{-i\boldsymbol{k}\cdot\boldsymbol{x}} \sqrt{\hat{c}(\boldsymbol{k})} \hat{n}_{\boldsymbol{k}} d\boldsymbol{k}$$

with  $\hat{n}_{k}$  complex white noise, i.e.:

 $\hat{n}_{k}$  complex-valued, Gaussian,  $\hat{n}_{-k} = \overline{\hat{n}_{k}}$ ,  $\mathbb{E}[\hat{n}_{k}] = 0$ , and  $\mathbb{E}[\hat{n}_{k}\overline{\hat{n}_{k'}}] = (2\pi)^{d}\delta(k - k')$ . (the representation is formal, one should use stochastic integrals  $d\hat{W}_{k} = \hat{n}_{k}dk$ ).

We have  $\hat{n}_{k} = \int e^{i \mathbf{k} \cdot \mathbf{x}} n(\mathbf{x}) d\mathbf{x}$  where  $n(\mathbf{x})$  is a real white noise, i.e.:  $n(\mathbf{x})$  real-valued, Gaussian,  $\mathbb{E}[n(\mathbf{x})] = 0$ , and  $\mathbb{E}[n(\mathbf{x})n(\mathbf{x}')] = \delta(\mathbf{x} - \mathbf{x}')$ . (in 1D, formally,  $n(x) = dW_x/dx$ ).

• Simulation (d = 1): in order to simulate  $(\mu(x_1), \ldots, \mu(x_n)), x_j = (j-1)h$ :

- compute the covariance vector  $\boldsymbol{C} = (c(x_1), \ldots, c(x_n)).$ 

- generate a random vector  $\mathbf{X} = (X_1, \ldots, X_n)$  of *n* independent Gaussian random variables with mean 0 and variance 1.

- filter with the square root of the Fourier transform of C:

$$\boldsymbol{Y} = \mathrm{IDFT} \left( \sqrt{\mathrm{DFT}(\boldsymbol{C})} \times \mathrm{DFT}(\boldsymbol{X}) \right)$$

 $\hookrightarrow \mathbf{Y}$  is a realization of  $(\mu(x_1), \ldots, \mu(x_n))$  (in practice, use FFT and IFFT).

## Random differential equations and ordinary differential equations

Goal: determine the limit  $\lim_{\varepsilon \to 0} X^{\varepsilon}(z)$  where

$$\frac{dX^{\varepsilon}}{dz} = F\left(\frac{z}{\varepsilon}\right)$$

for a fairly general random process  $(F(z))_{z\geq 0}$ .

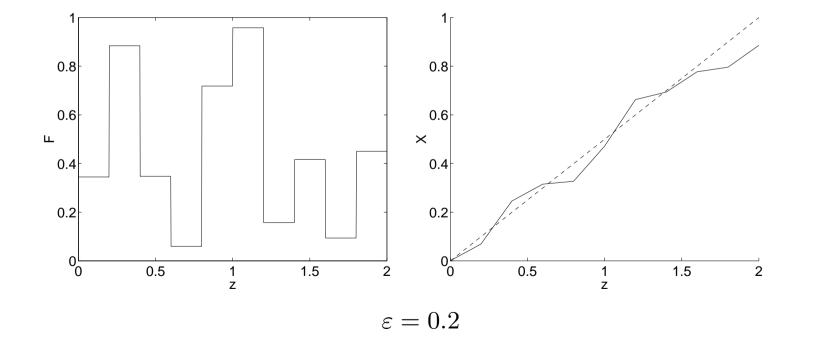
### Method of averaging: Toy model

Let  $X^{\varepsilon}(z) \in \mathbb{R}$  be the solution of

$$\frac{dX^{\varepsilon}}{dz} = F(\frac{z}{\varepsilon})$$

with  $F(z) = \sum_{i=1}^{\infty} F_i \mathbf{1}_{[i-1,i)}(z)$ ,  $F_i$  independent random variables  $\mathbb{E}[F_i] = \overline{F}$  and  $\mathbb{E}[(F_i - \overline{F})^2] = \sigma^2$ .

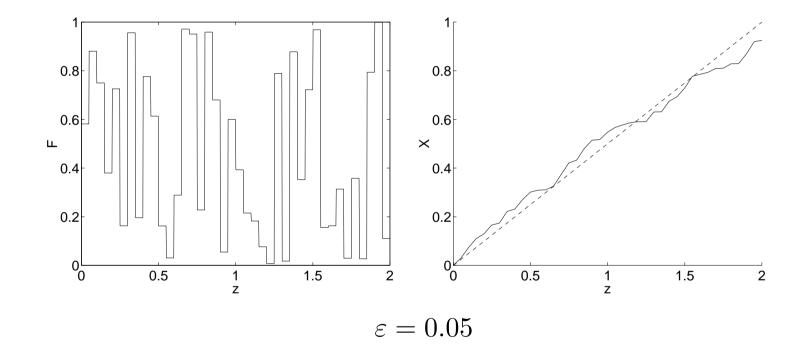
 $(z \mapsto t, \text{ particle in a random velocity field})$ 

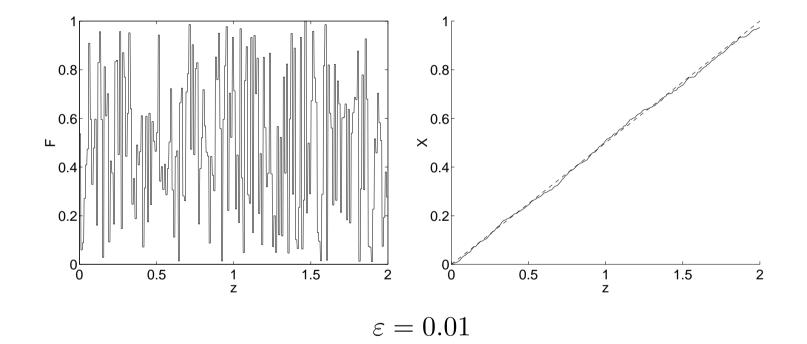


$$\begin{split} X^{\varepsilon}(z) &= \int_{0}^{z} F\left(\frac{s}{\varepsilon}\right) ds = \varepsilon \int_{0}^{\frac{z}{\varepsilon}} F(s) ds = \varepsilon \left(\sum_{i=1}^{\left\lfloor \frac{z}{\varepsilon} \right\rfloor} F_{i}\right) + \varepsilon \int_{\left\lfloor \frac{z}{\varepsilon} \right\rfloor}^{\frac{z}{\varepsilon}} F(s) ds \\ &= \varepsilon \left\lfloor \frac{z}{\varepsilon} \right\rfloor \times \frac{1}{\left\lfloor \frac{z}{\varepsilon} \right\rfloor} \left(\sum_{i=1}^{\left\lfloor \frac{z}{\varepsilon} \right\rfloor} F_{i}\right) + \varepsilon \left(\frac{z}{\varepsilon} - \left\lfloor \frac{z}{\varepsilon} \right\rfloor\right) F_{\left\lfloor \frac{z}{\varepsilon} \right\rfloor} \\ &\varepsilon \to 0 \downarrow \\ z & \text{a.s. } \downarrow (LLN) \\ \mathbb{E}[F_{1}] = \bar{F} \end{split}$$

Thus:

$$X^{\varepsilon}(z) \xrightarrow{\varepsilon \to 0} \bar{X}(z), \qquad \frac{d\bar{X}}{dz} = \bar{F}.$$





Goal: determine the limit  $\lim_{\varepsilon \to 0} X^{\varepsilon}(z)$  where

$$\frac{dX^{\varepsilon}}{dz} = F\left(\frac{z}{\varepsilon}\right)$$

for a stationary random process  $(F(z))_{z\geq 0}$ .

The previous analysis can be extended provided

$$\frac{1}{Z} \int_0^Z F(z) dz \xrightarrow{Z \to \infty} \bar{F}$$

(i.e. F is ergodic)

Next goal: determine the limit  $\lim_{\varepsilon \to 0} X^{\varepsilon}(z)$  where

$$\frac{d\boldsymbol{X}^{\varepsilon}}{dz} = \boldsymbol{F}\left(\frac{z}{\varepsilon}, \boldsymbol{X}^{\varepsilon}(z)\right)$$

### Method of averaging: Khasminskii theorem

$$\frac{d\boldsymbol{X}^{\varepsilon}}{dz} = \boldsymbol{F}(\frac{z}{\varepsilon}, \boldsymbol{X}^{\varepsilon}), \quad \boldsymbol{X}^{\varepsilon}(0) = \boldsymbol{x}_{0}$$

Assume:

 $\boldsymbol{x} \mapsto \boldsymbol{F}(z, \boldsymbol{x})$  is Lipschitz,  $z \mapsto \boldsymbol{F}(z, \boldsymbol{x})$  is stationary and ergodic. Define:

$$ar{m{F}}(m{x}) = \mathbb{E}[m{F}(z,m{x})]$$

Let  $\bar{X}$  be the solution of

$$\frac{d\boldsymbol{X}}{dz} = \bar{\boldsymbol{F}}(\bar{\boldsymbol{X}}), \quad \bar{\boldsymbol{X}}(0) = \boldsymbol{x}_0$$

Theorem: for any Z > 0,

$$\sup_{z \in [0,Z]} \mathbb{E} \left[ |\boldsymbol{X}^{\varepsilon}(z) - \bar{\boldsymbol{X}}(z)| \right] \stackrel{\varepsilon \to 0}{\longrightarrow} 0$$

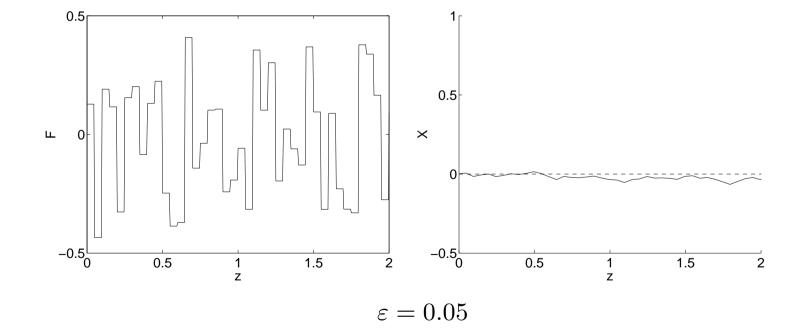
#### [1] R. Z. Khasminskii, Theory Probab. Appl. 11 (1966), 211-228.

## **Random differential equations and Brownian motion**

### Toy model

 $\frac{dX^{\varepsilon}}{dz} = F(\frac{z}{\varepsilon})$ 

with  $F(z) = \sum_{i=1}^{\infty} F_i \mathbf{1}_{[i-1,i)}(z)$ ,  $F_i$  independent random variables  $\mathbb{E}[F_i] = \overline{F} = 0$  and  $\mathbb{E}[(F_i - \overline{F})^2] = \sigma^2$ .



For any  $z \in [0, Z]$ , we have

$$X^{\varepsilon}(z) \xrightarrow{\varepsilon \to 0} \bar{X}(z), \qquad \frac{dX}{dz} = \bar{F} = 0.$$

No macroscopic evolution is noticeable.

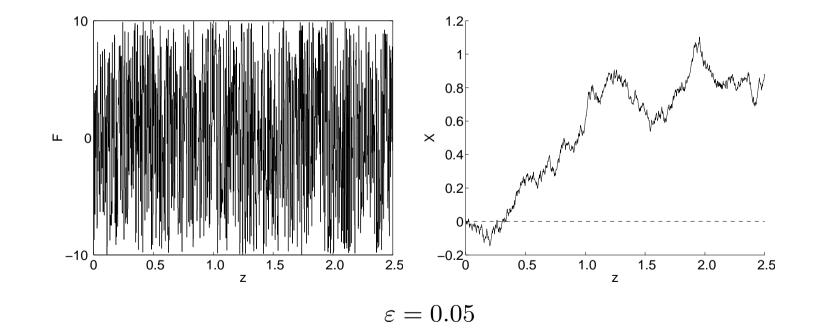
 $\rightarrow$  it is necessary to look at larger z to get an effective behavior

$$z \mapsto \frac{z}{\varepsilon}, \quad \tilde{X}^{\varepsilon}(z) = X^{\varepsilon}(\frac{z}{\varepsilon})$$
$$\frac{d\tilde{X}^{\varepsilon}}{dz} = \frac{1}{\varepsilon}F(\frac{z}{\varepsilon^2})$$

### **Diffusion-approximation:** Toy model

$$\frac{dX^{\varepsilon}}{dz} = \frac{1}{\varepsilon}F(\frac{z}{\varepsilon^2})$$

with  $F(z) = \sum_{i=1}^{\infty} F_i \mathbf{1}_{[i-1,i)}(z)$ ,  $F_i$  independent random variables  $\mathbb{E}[F_i] = 0$  and  $\mathbb{E}[F_i^2] = \sigma^2$ .



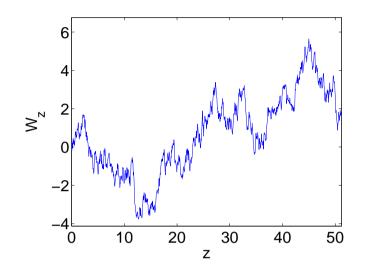
$$\begin{split} X^{\varepsilon}(z) &= \int_{0}^{z} \frac{1}{\varepsilon} F(\frac{s}{\varepsilon^{2}}) ds = \varepsilon \int_{0}^{\frac{z}{\varepsilon^{2}}} F(s) ds = \varepsilon \begin{pmatrix} \left[\frac{z}{\varepsilon^{2}}\right] \\ \sum_{i=1}^{z} F_{i} \end{pmatrix} + \varepsilon \int_{\left[\frac{z}{\varepsilon^{2}}\right]}^{\frac{z}{\varepsilon^{2}}} F(s) ds \\ &= \varepsilon \sqrt{\left[\frac{z}{\varepsilon^{2}}\right]} \times \frac{1}{\sqrt{\left[\frac{z}{\varepsilon^{2}}\right]}} \begin{pmatrix} \left[\frac{z}{\varepsilon^{2}}\right] \\ \sum_{i=1}^{z} F_{i} \end{pmatrix} + \varepsilon \begin{pmatrix} \frac{z}{\varepsilon^{2}} - \left[\frac{z}{\varepsilon^{2}}\right] \end{pmatrix} F_{\left[\frac{z}{\varepsilon^{2}}\right]} \\ &\varepsilon \to 0 \downarrow \\ \sqrt{z} & \text{law} \downarrow (CLT) \\ &\mathcal{N}(0, \sigma^{2}) & 0 \end{split}$$

Thus:  $X^{\varepsilon}(z)$  converges in distribution as  $\varepsilon \to 0$  to  $\bar{X}(z)$  whose distribution is  $\mathcal{N}(0, \sigma^2 z)$ .

With some more work: The process  $(X^{\varepsilon}(z))_{z\geq 0}$  converges in distribution to a Brownian motion  $(\sigma W_z)_{z\geq 0}$ .

# **Diffusion processes and stochastic differential equations**

## **Example: Brownian motion**



 $W_z$  (issued from 0): zero-mean Gaussian process with covariance  $\mathbb{E}[W_z W_{z'}] = z \wedge z'$ 

Its increments are independent and:

$$\mathbb{E}[(W_{z+h} - W_z)^2] = h$$

Let  $\phi$  be a bounded real function:

$$u(z,x) := \mathbb{E}[\phi(x+W_z)] = \int \frac{1}{\sqrt{2\pi z}} \exp\left(-\frac{w^2}{2z}\right) \phi(x+w) dw$$
$$= \int \underbrace{\frac{1}{\sqrt{2\pi z}} \exp\left(-\frac{(y-x)^2}{2z}\right)}_{p_z(x,y)} \phi(y) dy$$

For x, z fixed,  $y \mapsto p_z(x, y)$  is the pdf of the random variable  $x + W_z$ .  $p_z(x, y)$  is the kernel of the heat operator.

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### **Example: Brownian motion**

• The moment

$$u(z,x) := \mathbb{E}[\phi(x+W_z)] = \int p_z(x,y)\phi(y)dy, \qquad p_z(x,y) = \frac{1}{\sqrt{2\pi z}} \exp\left(-\frac{(y-x)^2}{2z}\right)$$

satisfies the (backward) Kolmogorov equation

$$\frac{\partial u}{\partial z} = \frac{1}{2} \frac{\partial^2 u}{\partial x^2}, \qquad u(z=0,x) = \phi(x)$$

Reciprocal: The partial differential equation

$$\frac{\partial u}{\partial z} = \frac{1}{2} \frac{\partial^2 u}{\partial x^2}, \qquad u(z=0,x) = \phi(x)$$

has the probabilistic representation:  $u(z,x) = \mathbb{E}[\phi(x+W_z)].$ 

The reciprocal is useful for Monte Carlo simulation techniques for solving PDEs.

• The pdf  $y \mapsto p_z(x, y)$  of  $x + W_z$  satisfies the Fokker-Planck (Kolmogorov forward) equation as a function of z and y:

$$\frac{\partial p_z}{\partial z} = \frac{1}{2} \frac{\partial^2 p_z}{\partial y^2}, \qquad p_{z=0}(x,y) = \delta(y-x)$$

#### **Stochastic differential equations**

Let X(z) be the solution of the one-dimensional stochastic differential equation

$$X(z) = x + \int_0^z \sigma(X(s))dW_s + \int_0^z b(X(s))ds$$

Existence and uniqueness ensured provided b and  $\sigma$  are  $\mathcal{C}^1$  with bounded derivatives. X(z) is continuous a.s., is squared integrable, is adapted (depends on  $(W_s)_{0 \le s \le z}$ ).

• Itô integral

$$\int_0^z \phi(X(s)) dW_s = \lim_{\Delta z \to 0} \sum_k \phi(X(z_k)) (W_{z_{k+1}} - W_{z_k}), \qquad 0 = z_0 < z_1 < \dots < z_n = z$$

Good properties: martingale (mean zero), Itô's isometry:

$$\mathbb{E}\left[\left(\int_0^z \phi(X(s))dW_s\right)^2\right] = \int_0^z \mathbb{E}[\phi(X(s))^2]ds$$

• Stratonovich integral:

$$\int_0^z \phi(X(s)) \circ dW_s = \lim_{\Delta z \to 0} \sum_k \frac{\phi(X(z_{k+1})) + \phi(X(z_k))}{2} (W_{z_{k+1}} - W_{z_k})$$

• Relation:

$$\int_{0}^{z} \phi(X(s)) \circ dW_{s} = \int_{0}^{z} \phi(X(s)) dW_{s} + \frac{1}{2} \int_{0}^{z} \sigma(X(s)) \phi'(X(s)) ds$$

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$$X(z) = x + \int_0^z \sigma(X(s))dW_s + \int_0^z b(X(s))ds$$

• Itô's formula:

$$\phi(X(z)) = \phi(x) + \int_0^z \sigma(X(s))\phi'(X(s))dW_s + \int_0^z b(X(s))\phi'(X(s))ds + \frac{1}{2}\int_0^z \sigma(X(s))^2\phi''(X(s))ds$$

• X(z) is a diffusion process with the generator  $Q = \frac{1}{2}\sigma^2(x)\frac{\partial^2}{\partial x^2} + b(x)\frac{\partial}{\partial x}$ 

• The moment  $u(z,x) := \mathbb{E}[\phi(X(z))|X(0) = x]$  satisfies the Kolmogorov equation:

$$\frac{\partial u}{\partial z} = Qu, \qquad u(z=0,x) = \phi(x)$$

• The pdf  $y \mapsto p_z(x, y)$  of X(z) (starting from X(0) = x) satisfies the Fokker-Planck equation as a function of z and y:

$$\frac{\partial p_z}{\partial z} = Q^* p_z, \qquad p_{z=0}(x,y) = \delta(x-y)$$

where  $Q^*$  is the adjoint operator of Q:

$$Q^*p(y) = \frac{1}{2} \frac{\partial^2}{\partial y^2} \left( \sigma^2(y) p(y) \right) - \frac{\partial}{\partial y} \left( b(y) p(y) \right)$$

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#### **Diffusion processes**

• Let  $\sigma$  and b be  $\mathcal{C}^1(\mathbb{R}, \mathbb{R})$  functions with bounded derivatives. Let  $W_z$  be a Brownian motion.

The solution X(z) of the 1D stochastic differential equation:

 $dX(z) = \sigma(X(z))dW_z + b(X(z))dz$ 

is a diffusion process with the generator

$$Q = \frac{1}{2}\sigma^2(x)\frac{\partial^2}{\partial x^2} + b(x)\frac{\partial}{\partial x}$$

• Let  $\boldsymbol{\sigma} \in \mathcal{C}^1(\mathbb{R}^n, \mathbb{R}^{n \times m})$  and  $\boldsymbol{b} \in \mathcal{C}^1(\mathbb{R}^n, \mathbb{R}^n)$  with bounded derivatives. Let  $\boldsymbol{W}_z$  be a *m*-dimensional Brownian motion. The solution  $\boldsymbol{X}(z)$  of the stochastic differential equation:

$$d\boldsymbol{X}(z) = \boldsymbol{\sigma}(\boldsymbol{X}(z))d\boldsymbol{W}_z + \boldsymbol{b}(\boldsymbol{X}(z))dz$$

is a diffusion process with the generator

$$Q = \frac{1}{2} \sum_{i,j=1}^{n} a_{ij}(\boldsymbol{x}) \frac{\partial^2}{\partial x_i \partial x_j} + \sum_{i=1}^{n} b_i(\boldsymbol{x}) \frac{\partial}{\partial x_i}$$

with  $\mathbf{a} = \boldsymbol{\sigma} \boldsymbol{\sigma}^T$ .

ESNT

## Random differential equations and stochastic differential equations

Next goal: determine the limit  $\lim_{\varepsilon \to 0} X^{\varepsilon}(z)$  where

$$\frac{dX^{\varepsilon}}{dz} = \frac{1}{\varepsilon} F\left(\frac{z}{\varepsilon^2}\right)$$

for a fairly general process  $(F(z))_{z\geq 0}$  with  $\mathbb{E}[F(z)] = 0$ .

Write

$$X^{\varepsilon}(z) = \frac{1}{\varepsilon} \int_{0}^{z} F\left(\frac{s}{\varepsilon^{2}}\right) ds = \varepsilon \sqrt{\frac{z}{\varepsilon^{2}}} \times \frac{1}{\sqrt{\frac{z}{\varepsilon^{2}}}} \int_{0}^{\frac{z}{\varepsilon^{2}}} F(s) ds$$

We know that

$$\begin{split} \mathbb{E}\Big[\frac{1}{\sqrt{\frac{z}{\varepsilon^2}}} \int_0^{\frac{z}{\varepsilon^2}} F\left(s\right) ds\Big] &= 0,\\ \lim_{\varepsilon \to 0} \mathbb{E}\Big[\Big(\frac{1}{\sqrt{\frac{z}{\varepsilon^2}}} \int_0^{\frac{z}{\varepsilon^2}} F\left(s\right) ds\Big)^2\Big] &= \lim_{Z \to \infty} Z\mathbb{E}\left[\Big(\frac{1}{Z} \int_0^Z F(s) ds\Big)^2\right]\\ &= 2\int_0^{\infty} \mathbb{E}[F(0)F(u)] du, \end{split}$$

The Gaussian property of the limit of  $X^{\varepsilon}(z)$  is ensured by an invariance principle. Conclusion:

$$\frac{1}{\varepsilon} \int_0^z F\left(\frac{s}{\varepsilon^2}\right) ds \xrightarrow{\varepsilon \to 0} \sqrt{2}\sigma W_z$$

in distribution, where  $(W_z)_{z\geq 0}$  is a Brownian motion and

$$\sigma^2 = \int_0^\infty \mathbb{E}[F(0)F(u)]du$$

Next goal: determine the limit  $\lim_{\varepsilon \to 0} X^{\varepsilon}(z)$  where

$$\frac{d\boldsymbol{X}^{\varepsilon}}{dz} = \frac{1}{\varepsilon} \boldsymbol{F}\left(\frac{z}{\varepsilon^2}, \boldsymbol{X}^{\varepsilon}(z)\right)$$

when

$$\mathbb{E}[\boldsymbol{F}(z, \boldsymbol{x})] = 0 \quad \forall \boldsymbol{x}$$

#### **Diffusion-approximation**

$$\frac{d\boldsymbol{X}^{\varepsilon}}{dz}(z) = \frac{1}{\varepsilon} \boldsymbol{F}\left(\frac{z}{\varepsilon^2}, \boldsymbol{X}^{\varepsilon}(z)\right), \qquad \boldsymbol{X}^{\varepsilon}(0) = \boldsymbol{x}_0 \in \mathbb{R}^d.$$

 $\boldsymbol{F}$  stationary, centered, and ergodic (in z):  $\mathbb{E}[\boldsymbol{F}(z, \boldsymbol{x})] = \boldsymbol{0}$ .

**Theorem:** The processes  $(\mathbf{X}^{\varepsilon}(z))_{z\geq 0}$  converge in distribution in  $\mathcal{C}^{0}([0,\infty), \mathbb{R}^{d})$  to the diffusion process  $\mathbf{X}$  with the generator  $\mathcal{L}$ :

$$\mathcal{L} = \sum_{i,j=1}^{d} a_{ij}(\boldsymbol{x}) \frac{\partial^2}{\partial x_i \partial x_j} + \sum_{j=1}^{d} b_j(\boldsymbol{x}) \frac{\partial}{\partial x_j}$$

with

$$a_{ij}(\boldsymbol{x}) = \int_0^\infty \mathbb{E} \left[ F_i(0, \boldsymbol{x}) F_j(u, \boldsymbol{x}) \right] du$$
  
$$b_j(\boldsymbol{x}) = \sum_{i=1}^d \int_0^\infty \mathbb{E} \left[ F_i(0, \boldsymbol{x}) \partial_{x_i} F_j(u, \boldsymbol{x}) \right] du$$

It means that the pdf  $p_z(\boldsymbol{x})$  of  $\boldsymbol{X}(z)$  satisfies:

$$\frac{\partial p_z}{\partial z} = \sum_{i,j=1}^d \frac{\partial^2}{\partial x_i \partial x_j} (a_{ij}(\boldsymbol{x}) p_z) - \sum_{j=1}^d \frac{\partial}{\partial x_j} (b_j(\boldsymbol{x}) p_z), \qquad p_{z=0}(\boldsymbol{x}) = \delta(\boldsymbol{x} - \boldsymbol{x}_0)$$

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#### **Diffusion-approximation - the one-dimensional case**

$$\frac{dX^{\varepsilon}}{dz}(z) = \frac{1}{\varepsilon} f\left(\frac{z}{\varepsilon^2}\right) h\left(X^{\varepsilon}(z)\right), \qquad X^{\varepsilon}(0) = x_0 \in \mathbb{R}.$$

f stationary, centered, and ergodic:  $\mathbb{E}[f(z)] = 0$ .

**Theorem:** The processes  $(X^{\varepsilon}(z))_{z\geq 0}$  converge in distribution in  $\mathcal{C}^{0}([0,\infty),\mathbb{R})$  to the diffusion process X with the generator  $\mathcal{L}$ :

$$\mathcal{L} = a(x)\frac{\partial^2}{\partial x^2} + b(x)\frac{\partial}{\partial x}$$

with

$$a(x) = \frac{1}{2}\sigma^2 h^2(x), \qquad b(x) = \frac{1}{2}\sigma^2 h(x)h'(x), \qquad \sigma^2 = 2\int_0^\infty \mathbb{E}\left[f(0)f(u)\right] du$$

It means that X(z) is the solution of the SDE

$$dX(z) = \sigma h(X(z))dW_z + b(X(z))dz$$

or

$$dX(z) = \sigma h(X(z)) \circ dW_z$$

Remember that  $\int_0^z \frac{1}{\varepsilon} f\left(\frac{s}{\varepsilon^2}\right) ds$  converges in distribution to  $\sigma W_z$ .  $\hookrightarrow$  the "natural" integral is Stratonovich.